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The role of aspect ratio and beta in H-mode confinement scalings

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Abstract

The addition of high power, low aspect ratio data from the NSTX and MAST experiments has motivated a new investigation of the effect of aspect ratio on confinement scaling. Various statistical methods, including those that incorporate estimates of measurement error, have been applied to datasets constrained by the standard set of criteria in addition to the range of κ and $M_{\rm eff}$ appropriate to ITER operation. Development of scalings using engineering parameters as predictor variables results in ε -scaling coefficients that range from 0.38 to 1.29; the transformation of these scalings to physics variables results in an unfavourable dependence of $B\tau$ on β , but a favourable dependence on ε . Because the low aspect ratio devices operate at low B_T and therefore high β_T ,

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a strong correlation exists between ε and β , and this makes scalings based on physics variables imprecise.

1. Introduction

Recent analysis of the international H-mode confinement database focused on extracting the dependence of the thermal confinement ($\tau_{E,\text{th}}$) on beta (β) and collisionality (ν_*) through the use of a variety of statistical methods [1]. In particular, it was shown that in an extreme case, the data could be consistent with a null dependence on β , in line with results from dedicated scans on DIII-D and JET [2,3], rather than a strongly unfavourable dependence such as in the IPB98(y, 2) scaling [4]. One of the important aspects of this work was the use of statistical methods in which the measurement error of the predictor ('independent') and response (dependent) variables could be included in the analysis. The use of this approach coupled with the more standard ones led to a range of parametric dependences that fit the data well.

The objective of the present paper is to study the role of aspect ratio (R/a) and beta on confinement by making use of the high power, low aspect ratio data recently contributed to the database from the NSTX and MAST devices. While some of this data was in the analysis in [1], there were too few observations to conduct a precise study of their effect on the scalings. Both NSTX and MAST operate at aspect ratios of R/a = 1.3 to 1.5, a factor of approximately two lower than that of ITER and most of the data in the H-mode database. In addition, both devices operate in a range of elongation (1.6 to 2.4) that encompasses the ITER target operating point, making the study of the aspect ratio effect relevant to the ITER shape. Prior to the inclusion of the low aspect ratio data, the dependence of confinement with this parameter was obtained from a dataset extending only to higher aspect ratio (up to 5.5), but with this high aspect ratio data provided by devices with cross-sectional shapes dissimilar from that of ITER. Both PDX and ASDEX operated near R/a = 4, but with circular cross-sections, and PBX-M ran at R/a = 5.5, but with indented, or 'bean-shaped', plasmas.

In this work, the role of aspect ratio in governing confinement trends will be examined by comparing results from a variety of statistical methods as applied to datasets with different constraints. These methods include ordinary least squares regression (OLSR), a Bayesian approach [5, 6] and the principal component error-in-variable (PCEIV) technique used in recent analyses [1, 7]. The latter two techniques take measurement error of the variables into account. Because the low aspect ratio devices operated at vacuum toroidal fields (B_T) that are an order of magnitude lower than those of conventional aspect devices, and therefore higher $\beta_T (= \langle p \rangle / (B_T^2/2\mu_0))$, a significant correlation exists between β and ε (the inverse aspect ratio $\varepsilon = a/R$, will be used as a predictor variable rather than aspect ratio itself). This makes a precise determination of the ε and β dependence difficult.

In section 2, a description of the new data and analysis selection criteria, along with the statistics and condition of the selected datasets, will be given. The results of the statistical analysis will be given in section 3, and an extrapolation of the results to ITER and to a component test facility (CTF) based on a low aspect ratio design [8] will be given in section 4.

2. Data and selection criteria

2.1. NSTX

NSTX operates with minor radius a = 0.67 m and major radius R = 0.85 m (R/a = 1.27). Over 90 H-mode discharges from the 2002 to 2004 experimental campaigns were contributed

to the international H-mode confinement database. The discharges were sawtooth-free, they were both ELM-free and ELMy with either Type I or smaller ELMs, and they exhibited periods of stationary stored energies for at least several energy confinement times. The parameter ranges of the data are plasma current $I_{\rm p}=0.6$ –1.2 MA, $B_{\rm T}=0.3$ –0.5 T, line-averaged density $n_{\rm e}=(1.5$ –7) \times 10^{19} m⁻³ and elongation $\kappa=1.7$ –2.4, and with D⁰ neutral beam injection, with thermal power loss $P_{\rm L,th}(P_{\rm heat}-{\rm d}W/{\rm d}t)$ from 1.9 to 7.1 MW at beam energies up to 100 keV, into D⁺ plasmas. Plasma operation was in either double null (DN) or lower single null (SNL) divertor configurations; the SNL configuration had the ion grad-B drift in the favourable direction. The data exhibited enhancement factors of 0.6–1.4 relative to the IPB98(y, 2) scaling, and while the data scaled nearly linearly with $I_{\rm p}$ at fixed $B_{\rm T}$, the data statistically exhibited a weaker $I_{\rm p}$ and stronger $B_{\rm T}$ dependence than that given by this scaling, with $\tau_{E,th} \sim I_{\rm p}^{0.5} B_{\rm T}^{1}$ [9, 10].

2.2. MAST

The MAST data represent quasi-stationary ELMy plasmas with and without sawteeth. The dataset covers the engineering parameter range $I_{\rm p}=0.73$ –0.78 MA, $R_{\rm geo}=0.8$ –0.83 m, a=0.54–0.57 m (R/a=1.47), $\kappa=1.9$ –2.0, triangularity $\delta=0.44$ –0.52, $B_{\rm T}=0.45$ –0.49 T, $n_{\rm e}=(3.0$ –5.4) \times 10¹⁹ m⁻³ and $P_{\rm L,th}=1.5$ –2.4 MW. All data were from a DN divertor configuration. The working gas was deuterium and plasmas were heated with neutral beams with energy 40 keV. Typically, one of the two beam lines was operated in hydrogen to improve the ion temperature measurement, but the amount of hydrogen from the single beam was small, and effective mass $M_{\rm eff}=1.94$ –2.0, as confirmed by neutral particle analyser measurements. The effective charge measured by bremsstrahlung emission at mid-radius was $Z_{\rm eff}=1.1$ –1.4. The thermal energy confinement time on MAST broadly agrees with the IPB98(y, 2) scaling. It is observed, however, that the plasmas with lower collisionality have better confinement normalized to IPB98(y, 2) than high collisionality plasmas [11].

2.3. Data selection

The data to be used for this study come from the ITPA confinement database DB4V2, which is based on the now public DB3V13 but with the addition of 92 NSTX, 7 MAST and 507 JET datapoints. The data selected for this study were constrained first by the standard selection, which, among other criteria, is based on limiting fast ion content and ensuring near-stationary conditions ([1] and references therein). Furthermore, for this study, whose aim is to isolate the effect of aspect ratio, the data were constrained to both elongation and $M_{\rm eff}$ ranges from 1.6 to 2.4. The effect of shaping on the β dependence is studied in [12]. The number of observations satisfying these constraints for each tokamak are ASDEX-U (509), C-Mod (31), COMPASS (16), DIII-D (264), JET (1487), MAST (9), NSTX (53), PBX-M (36) and START (7). The weighting of the data for each tokamak to be used in the subsequent statistical analyses is $w_j = 1/N_j^{1/2}$, where N_j is the number of observations for the jth tokamak. This weighting is different from the $1/(2 + N_j^{1/2}/4)$ weighting used in [1] and references therein, in which the effect from tokamaks with the greatest number of observations still dominate. At another extreme, a $1/N_i$ weighting would treat the data from each machine equally. While we adopt the compromise $1/N_i^{1/2}$ weighting, results from the other two weightings will be discussed. Measurement errors for the engineering parameters are as in [1,7]: $\delta I_p = 1.3\%$, $\delta B_T = 1.5\%$, $\delta n_{\rm e} = 5\%$, $\delta P = 14\%$, $\delta R = 1.3\%$ and $\delta \varepsilon = 3.2\%$.

The addition of the low aspect ratio data in the ITPA database extends the range of both β and ε significantly. This is shown in figure 1, where $\beta_{\text{th,TOT}}$ is plotted as a function of ε for the

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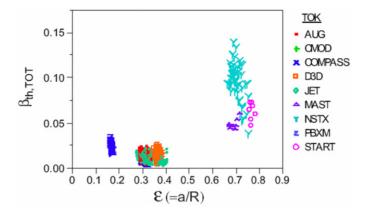


Figure 1. Range of $\beta_{\text{th.TOT}}$ and ε (inverse aspect ratio) from the constrained ITPA H-mode database.

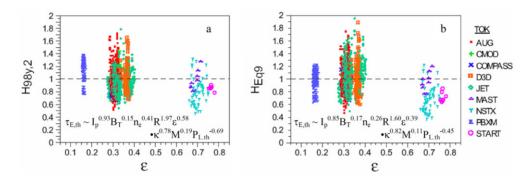


Figure 2. Confinement enhancement factor as a function of inverse aspect ratio for the IPB98(y, 2) scaling (a) and from equation (9) in [1] (b).

set of data satisfying the above constraints. Here, $\beta_{\text{th,TOT}}$ is the thermal pressure normalized to the total magnetic field, defined as $B_{\text{TOT}} = [B_{\text{T}}^2 + B_{\text{pol}}^2]^{1/2}$, where B_{T} is the vacuum toroidal field at the geometric axis and B_{pol} is the poloidal field determined from q_{95} . B_{TOT} is used in the calculation of the dimensionless parameters since, in a low aspect ratio device, $B_{\text{pol}} \sim B_{\text{tor}}$ at the plasma edge, and thus the poloidal field influences the plasma beta and normalized gyroradius $\rho_*(\sim T_{\text{e}}^{1/2}/\varepsilon R B_{\text{TOT}})$. Furthermore, the use of the total magnetic field is consistent with the Troyon definition of β [13]. For the low aspect ratio devices, the difference between the toroidal, B_{T} , and the total magnetic field, B_{TOT} , can be up to 25%.

As can be seen in figure 1, the addition of MAST and NSTX extends the range of $\beta_{th,TOT}$ by a factor of approximately five, and the range of ε by over a factor of two. In principle, this should help improve the determination of the scaling with these parameters. However, as seen in the figure and as was pointed out in [11], the fact that the higher ε devices operate at higher β introduces a strong correlation between these two parameters. It is also seen that the lower ε PBX-M data is at higher $\beta_{th,TOT}$ as well, due to operation in its indented configuration. It will be seen that including PBX-M data in this analysis will reduce the favourable ε dependence for a given β dependence. The NSTX and MAST data lie in a collisionality regime that is well within the range of data at lower ε , but at 50% higher ρ_* due to operation at low B_T .

Examples of how well previously published thermal confinement scalings describe the ε dependence are shown in figure 2. Figure 2(a) shows the thermal confinement enhancement

0.01

-0.63

ln ε

 $ln I_p$ $ln B_T$ 0.50 0.31 -0.040.82 0.86 $\ln I_{\rm p}$ $\ln B_{\rm T}$ 0.49 0.43 0.48 0.40 -0.84-0.020.43 -0.33-0.01 -0.17 $\ln n_{\circ}$ 1 ln R0.74 0.48 -0.331 0.77 -0.590.83 0.41 0.00 0.77 1 -0.28In F

-0.12

Table 1. Pairwise correlations for the $\kappa = 1.6$ –2.4, $M_{\rm eff} = 1.6$ –2.4 dataset with (red) and without (black) PBX-M. The strongest correlations are denoted by the bold/italicized font.

factor H as a function of ε or the IPB98(y, 2) scaling, while figure 2(b) shows the H-factor relative to the scaling given in equation (9) of [1], where the confinement is taken to scale as β^0 . The IPB98(y, 2) scaling goes as $\varepsilon^{0.58}$ when expressed in engineering variables, while the scaling given in equation (9) of [1] goes as $\varepsilon^{0.39}$. In both the cases, the scalings overpredict the confinement time for high ε data (NSTX, MAST and START), as indicated by H < 1 for these scalings for those devices. Similar results are found for the other scalings presented in [1].

-0.48

-0.14

1

To perform the statistical analyses, two sets of predictor variables are chosen; engineering $(I_p, B_T, n_e, R, P_{L,th} \text{ and } \varepsilon)$ and physics $(\rho_*, \beta, \nu_*, q \text{ and } \varepsilon)$. B_T is used in the engineering variable fits since use of B_{TOT} would lead to a coupling of this parameter with I_p . The physics variables are defined with respect to volume-averaged values of thermal plasma density, total magnetic field and q_{cyl} . The engineering variables show some significant pairwise correlations among the variables. The most significant correlations in the dataset that includes PBX-M are between I_p/R , $I_p/P_{L,th}$ and $R/P_{L,th}$. In the dataset without PBX-M, the correlation between B_T/ε is also high, reflecting the correlation between β and ε seen in figure 1. The correlations for both datasets are given in table 1.

Using the physics variables as a predictor set introduces the issue that several of them are functions of the same variables, such as stored energy, ε and q (e.g. $\rho_* \sim W^{1/2}/\varepsilon$, $\beta \sim W$, $\nu_* \sim q/(\varepsilon^{3/2}W^2)$, $q \sim \varepsilon^2$). Consequently, correlations in the variables and their errors emerge in part due to these common factors. Using the definitions of the physics variables given in [4], significant correlations exist between ρ_*/ε , β/ε and ρ_*/β , and even in a 'reduced' set, with the common factor of stored energy normalized out [1], correlations still exist between ν_*/ρ , ν_*/β and ε/β .

3. Statistical results

Three statistical analysis methods have been used to analyse the confinement trends. OLSR is a standard technique that is used with the underlying assumption that the principal measurement error is in the response variable. The PCEIV method was employed in [1] to account for errors. In this approach the principal components (PC) of the logs of the predictor and response variables normalized by their errors are computed. A PC with a zero eigenvalue gives an exact linear relation among the variables. The minimum eigenvalue in practice is not zero, but it is small and is related to the scatter of the data about the PC. For this PC, then, the eigenvalue is assumed to be zero in order to determine this linear relation.

Finally, a Bayesian analysis is used [5,6]. Bayesian inference explicitly uses probability models to quantify data uncertainties and to fit the data. In this first application of this approach to this set of confinement data, we assume that the predictor variables are uncorrelated and that their actual value for each observation is assumed to be within a normal distribution

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Table 2a. $\tau_{E,\text{th}}$ scaling coefficients for the engineering predictor variables for three different statistical techniques. The uncertainties in the exponents are shown for the OLSR and Bayesian cases.

Case		Coefficients	$I_{\rm p}$	B_{T}	$n_{\rm e}$	R	$P_{ m L,th}$	ε	RMSE
1	OLSR	6.25e - 10	0.80	0.32	0.39	2.12	-0.66	0.95	0.166
			± 0.02	± 0.02	± 0.01	± 0.01	± 0.01	± 0.06	
2	PCEIV	1.02e - 9	0.66	0.42	0.42	2.28	-0.64	1.29	0.172
3	PCEIV-C	$9.43e{-10}$	0.73	0.36	0.39	2.14	-0.62	1.03	0.169
4	Bayesian	3.40e - 9	0.96	0.16	0.30	1.83	-0.67	0.38	0.160
			± 0.03	± 0.03	± 0.02	± 0.07	± 0.01	± 0.10	
5	OLSR-NBI only	4.78e - 9	0.83	0.29	0.34	2.05	-0.68	0.76	0.160
			± 0.02	± 0.02	± 0.01	± 0.04	± 0.01	± 0.06	

Table 2b. $B_{TOT}\tau_{E,th}$ scaling coefficients for the physics variables. The first four rows are the coefficients determined from the transformation of the results for the engineering variables given in table 2a, and the last three rows are the results using the reduced physics variables (i.e. stored energy normalized out of all except β) as predictor variables.

Case		Coefficients	$ ho_*$	β	ν_*	$q_{ m cyl}$	ε	RMSE
1a	OLSR		-2.86	-0.70	-0.09	-2.26	0.62	
2a	PCEIV		-2.93	-0.51	-0.10	-1.73	0.61	
3a	PCEIV-C		-2.78	-0.48	-0.12	-1.80	0.33	
4a	Bayesian		-2.67	-1.01	-0.11	-2.80	0.27	
5a	OLSR-NBI		-2.73	-0.96	-0.10	-2.49	0.44	
5	OLSR	2.40e - 7	-2.28	0.19	-0.36	-0.32	-1.68	0.308
6	PCEIV	$6.48e{-8}$	-2.72	-0.14	-0.18	-0.95	-0.37	0.396
7	Bayesian	$4.85e{-7}$	-2.16	0.17	-0.40	-0.39	-1.58	0.303

centred about the measurement. The width of the normal distribution is such that 99% of all possible values are contained within three standard deviations of the measured value. The three standard deviations constitute the measurement error. The value of the dependent variable is also assumed to be within a normal distribution, but with unknown error (i.e. large variance). A log-linear model was assumed for the fit, and the data were not weighted. Future studies using this method will focus on varying the models and the assumptions within the models.

In order to develop the range of scalings for discharge shapes similar to that of ITER, the dataset without PBX-M will be used as a base case. The OLSR for this dataset, with the datapoints weighted as described in section 2, yields the following relation:

$$\tau_{E,\text{th}} = 6.25e - 10I_{p}^{0.80}B_{T}^{0.32}n_{e}^{0.32}R_{L,\text{th}}^{-0.66}\varepsilon^{0.94}$$
(1)

in units of s, A, T, m⁻³, m and W. This is similar to the IPB98(y, 2) scaling but with a stronger ε dependence ($\tau \sim \varepsilon^{0.58}$ in IPB98(y, 2)). When transformed to physics variables, equation (1) can be expressed as

$$B_{\rm T}\tau \sim B_{\rm T}^{0.03} \rho_*^{-2.86} \beta_{\rm T}^{-0.70} \nu_*^{-0.09} q^{-2.26} \varepsilon^{0.62}$$
 (2)

exhibiting an unfavourable dependence on β and a weaker explicit dependence on ε than when expressed in terms of the engineering variables. The normalized confinement also exhibits a favourable dependence with decreasing q. The small residual dimensional dependence ($B^{0.03}$) indicates that the scaling nearly satisfies the Kadomtsev constraint (B^0 in this description). A comparison among the different methods is given in tables 2a and 2b. Table 2a shows the results based on the engineering predictor variables, while the first five rows of table 2b show these results when transformed to physics variables. The last three rows of table 2b show results

using the reduced physics variables as the predictor set. For all the cases except for Case 3, the measurement errors among the variables are assumed to be uncorrelated. Correlated errors do exist, however, for certain variables such as $\varepsilon(=a/R)$ and $B_T(=B_0R_0/R)$, where the subscript '0' refers to the value at the centre stack). These correlated errors can be taken into account using the PCEIV method, and the results based on them are given as PCEIV-C (Case 3) [14]. The root mean square error (RMSE) is defined as RMSE = $\sqrt{\sum_{i=1}^{N} (y_{\rm exp} - y_{\rm fit})^2/(N-2)}$, where $y_{\rm exp}$ is the natural log of the experimental value and $y_{\rm fit}$ is that of the value from the scaling.

The tables indicate a range of scalings and ε dependences, depending on which method is used. In table 2a, the ε dependence determined by the OLSR, PCEIV and PCEIV-C methods are stronger, and more favourable for low aspect ratio, than those in IPB98(y, 2) or the scalings determined in [1]. PCEIV-C gives a scaling with a slightly weaker ε dependence than PCEIV, and an interplay between I_p and B_T is also seen by comparing Cases 2 and 3. The Bayesian approach gives the weakest ε dependence and the lowest RMSE. This approach, however, does not weight the datapoints, and thus the result is biased towards the devices with the largest number of datapoints (JET, DIII-D and ASDEX-U). Indeed, when different weightings are used, different scaling results are obtained. As examples, using the $1/(2+N_j^{1/2}/4)$ weighting from [1], which does not de-emphasize the larger data contributions as much as $1/N_j^{1/2}$, OLSR yields

$$\tau_{E,\text{th}} \sim I_{\text{p}}^{0.85} B_{\text{T}}^{0.27} n_{\text{e}}^{0.35} R^{2.01} P_{\text{L,th}}^{-0.66} \varepsilon^{0.73},$$
(3)

while weighting each tokamak equally (i.e. through a $1/N_i$ weighting), OLSR gives

$$\tau_{E,\text{th}} \sim I_{\rm p}^{0.70} B_{\rm T}^{0.44} n_{\rm e}^{0.44} R^{2.32} P_{\rm L,th}^{-0.68} \varepsilon^{1.35}.$$
(4)

The influence of NSTX, with its weaker I_p dependence, is seen in equation (4) as the influence of this device is increased. Note also the stronger ε dependence. Case 5 is an OLSR result when the standard dataset is constrained to discharges that are heated only by neutral beam injection. The fit is similar to that in Case 1 within most exponent uncertainties, but with a significantly weaker dependence on ε . OLSR scalings with elongation included, where $\tau_{E,\text{th}}/\kappa^{\alpha}$, $\alpha=0.5$, 0.75 and 1.0, result in scalings with similar RMSEs as Case 1 and, as an example, for $\alpha=1.0$, the coefficients for $(I_p, B_T, n_e, R, P_{L,\text{th}} \text{ and } \varepsilon)$ become (0.76, 0.41, 0.41, 2.20, -0.71 and 1.06).

When transformed to physics variables, all the scalings based on the engineering variable predictor set show an unfavourable scaling on β . The ρ_* and ε dependences, when transformed from engineering variables, are sensitive to several of the engineering variable dependences, specifically those of n_e , R and $P_{L,th}$. Using the OLSR result (equation (1)) as an example, when the engineering coefficients are varied within their uncertainties (see table 1), the transformations to physics variables lead to the following ranges:

$$B\tau \sim \rho_*^{-(2.67-3.04)}\beta^{-(0.59-0.82)}\nu_*^{-(0.06-0.13)}q^{-(2.14-2.39)}\varepsilon^{(0.43-0.81)}.$$

The scalings based directly on the reduced physics parameter predictor set (Cases 5–7) indicate a weak dependence on β , but an unfavourable dependence on ε . The dependence on ν_* , however, is more favourable than that in the transformations from the engineering variable scalings (i.e. higher $B\tau$ for lower ν_*). The fits based on the physics parameters, however, are poor (in terms of the RMSE metric), due to correlations among the variables and their errors. The systematic differences in the β and ε dependences between the engineering and the reduced physics variable scalings result also from the differences in the dependent variable

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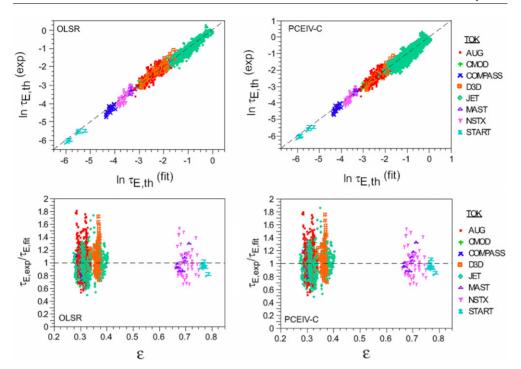


Figure 3. Scalings based on the engineering variable predictor set for OLSR (Case 1) and PCEIV-C (Case 3). Top panels show the experimental versus fitted confinement times, while the bottom panels show the confinement time enhancements versus ε .

used for these scalings (i.e. P instead of τ or W_{th}). This is discussed in more detail in [12]. These physics variable scalings, therefore, are shown for comparison purposes and will not be used for predictions. Fits based on the engineering variables (Cases 1 and 3) are shown in figure 3. Also shown is the $\tau_{E,th}^{\rm exp}/\tau_{E,th}^{\rm fit}$ as a function of ε for the scalings.

The influence of PBX-M can be assessed by determining the scaling with the data for this device included. Using the PCEIV method as an example, the resulting scaling is

$$\tau_{E,\text{th}} \sim I_{p}^{0.90} B_{T}^{0.13} n_{e}^{0.31} R^{1.77} P_{L,\text{th}}^{-0.55} \varepsilon^{0.21}$$
(5)

exhibiting a stronger I_p , weaker B_T and a much weaker ε dependence when compared with the results for Case 2. The latter is due to the relatively higher β and τ of PBX-M at the lowest ε , possibly due to its indented shape. We also note that excluding data from the low power, small, low aspect ratio device START has an insignificant impact on the scalings.

In addition to the range of scaling coefficients that result from applying different statistical techniques, the precision of the estimates with respect to the measurement errors can be assessed. An illustration of this is shown in figure 4, where the range of coefficients obtained when the errors on W_{th} and $P_{L,th}$ were varied by $\pm 50\%$ are plotted. The red shaded region shows the range of power and ε coefficients (α_P , α_ε), while the blue shaded region shows the range of β and ε coefficients resulting from the transformation from engineering to physics variables. The diamonds show the values of the coefficients for the experimental measurement error (14% for both W_{th} and $P_{L,th}$). Clearly seen is the interplay between the β and ε exponents, indicative of their correlation.

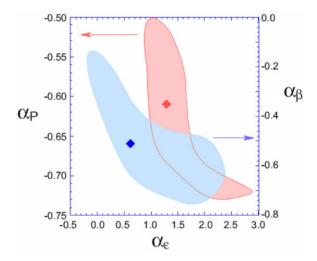


Figure 4. Range of exponents for power (red) and β (blue) as a function of ε from error analysis based on the PCEIV method. The $\beta - \varepsilon$ exponent $(\alpha_{\beta}, \alpha_{\varepsilon})$ range is transformed from the power- $\varepsilon(\alpha_{P}, \alpha_{\varepsilon})$ range.

4. Summary and predictions

The addition of high power, low aspect ratio data from the NSTX and MAST experiments has allowed an investigation of the effect of aspect ratio on confinement scaling. Various statistical methods, including those that incorporate estimates of measurement error, have been applied to datasets constrained by the standard set of criteria in addition to the range of κ and $M_{\rm eff}$ appropriate to ITER operation. Development of scalings using engineering variables as the predictor set results in ε -scaling coefficients that range from 0.38 to 1.29; the transformation of these scalings to physics variables results in unfavourable dependences on β . Because the low aspect ratio devices operate at low magnetic field and high $\beta_{\rm T}$, a strong correlation exists between ε and β , making scalings based on physics variables unreliable.

Not one scaling expression that can be identified as the 'best scaling' emerges from this analysis; consequently, the scalings developed here can be used as a set to estimate only a range of possible confinement times for future devices. This is done for both ITER and a low aspect ratio design for a CTF [8], and the results are given in table 3. Two confinement estimates are given for ITER: one for the base case ($\beta_T \sim 2.6\%$), where the loss power is held constant at 87 MW, and one for an enhanced performance case, where the loss power was adjusted in order for $\beta_t = 4\%$, as was done in [1]. No $M_{\rm eff}$ corrections were made to the estimates since the $M_{\rm eff}$ dependences, as determined in [1], were weak. Three methods (Cases 1, 2 and 3) show similar confinement predictions for the base case of ITER, while the scaling based on NB-only data (Case 5) is lower. The β -degradation in all these scalings causes much lower confinement estimates at $\beta_1 = 4\%$.

The confinement estimates for the CTF are based on the lowest divertor power flux case given in [8]. Because the CTF operates at higher elongation than ITER ($\kappa = 3.2$), a correction factor of $(3.2/2)^{0.75}$, consistent with the scalings developed in [1], was applied to the confinement estimates. For CTF, a confinement time of 0.32 s is required to meet its objectives. The confinement estimates based on Cases 1–3 indicate that a confinement enhancement factor of approximately two (over these H-mode scaling estimates) is necessary. CTF, however, may operate in a hot ion mode in which the ion transport would be close to neoclassical levels while

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Table 3. Energy confinement time predictions for ITER and CTF.

	Predictor				
Case	variables	Method	$(\beta_{\rm t} \sim 2.6\%)$	$(\beta_t = 4\%)$	CTF
1	Engineering	OLSR	3.47 s	1.33 s	0.16 s
2	Engineering	PCEIV	3.46 s	1.41 s	0.15 s
3	Engineering	PCEIV-C	3.46 s	1.56 s	0.16 s
5	Engineering	OLSR-NBI	3.16 s	0.89 s	$0.14 \mathrm{s}$

the electron transport would be dominant. In this case, the confinement time for CTF would be greater than those given by the predictions in table 3 [8].

Additional data from the low aspect ratio devices, including those from dedicated β -scans and from aspect ratio similarity experiments, will help further refine the β and ε dependences reported here. In particular, data from low aspect ratio plasmas with $\beta_{\text{th,TOT}}$ values that overlap those from conventional aspect ratio would help improve the condition of the database. This can be accomplished by running either low input power, low R/a plasmas or low B_T , high input power plasmas at higher aspect ratio.

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